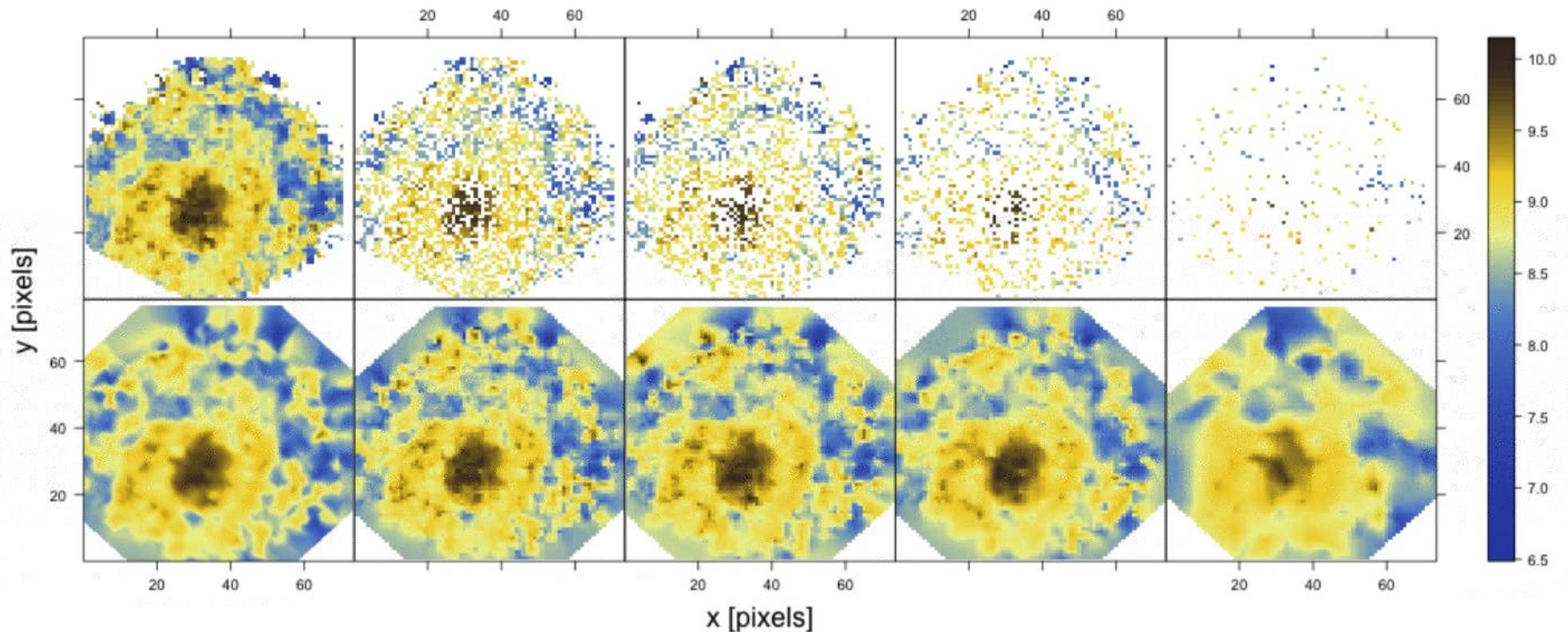


# Spatial Field Reconstruction with INLA

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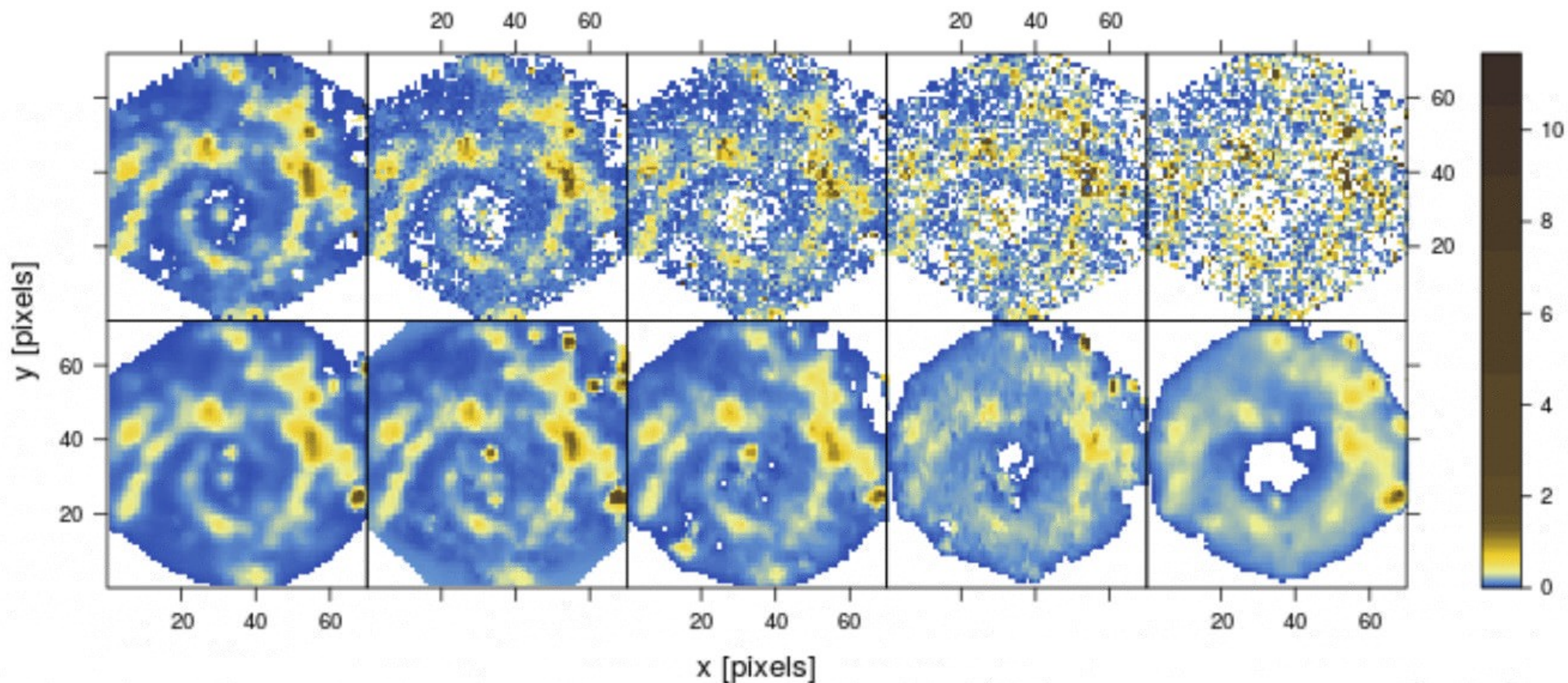
November 12th, 2019

# Examples



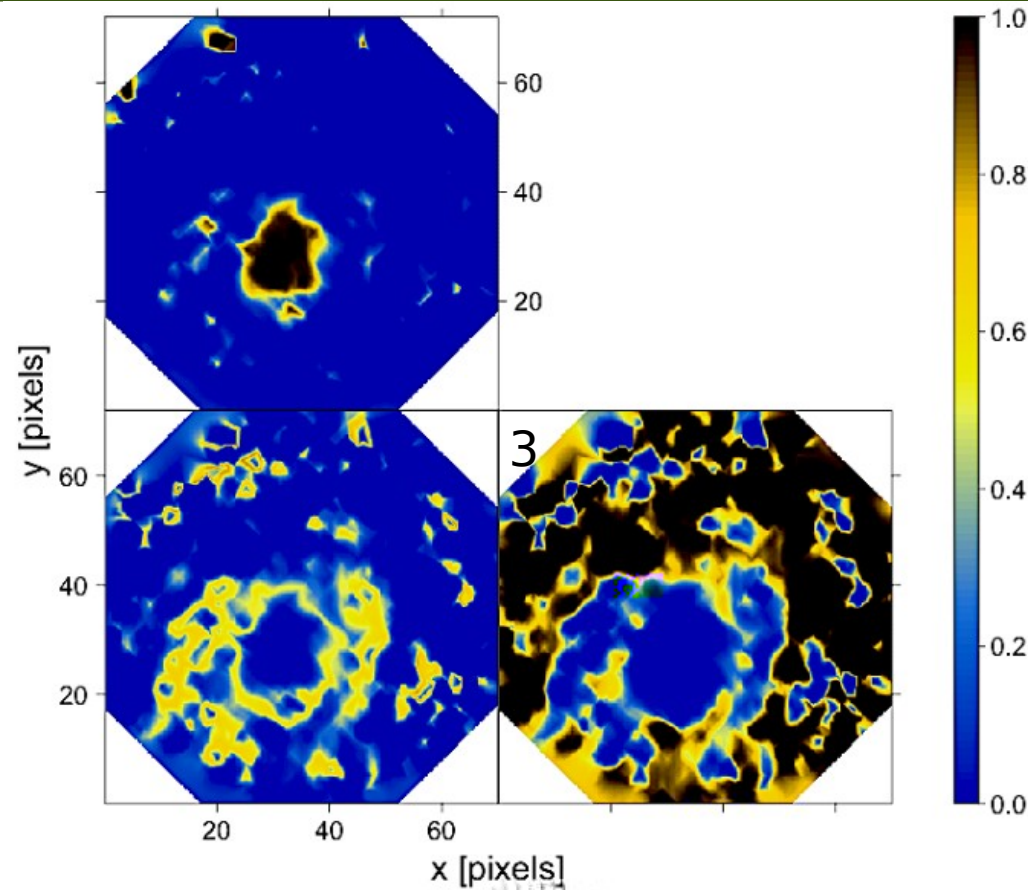
Predictions from INLA for input starlight age of NGC 0309 when 100, 75, 50, 25 and 5% (left to right) of the data is used. Upper panels show the starlight input, bottom the INLA prediction (source, González-Gaitán et al. (2018)).

# Examples



Predictions from INLA for H $\alpha$  EW map of NGC 0309 with S/N of 10, 2, 1, 0.5 and 0.3 (left to right). Upper panels show input and bottom panels INLA predictions (source, González-Gaitán et al. (2018)).

# Examples



Probability maps of NGC 0309 for three ranges of age,  $\log(t[\text{yr}])$ , arranged anticlockwise:  $\log(t[\text{yr}]) > 9.5$ ,  $9.3 > \log(t[\text{yr}]) > 9.0$  and  $\log(t[\text{yr}]) < 8.9$ . The bins were chosen to represent bottom (<2.5%), middle (32%–68%) and top (>97.5%) quantiles of the reconstructed population age map (source, González-Gaitán et al. (2018)).

# Structure

- **Introduction**
  - Bayesian Statistics
  - Latent Gaussian Models (LGMs)
    - Notation & Properties
  - Inference
- **INLA**
  - A different approach
  - The Method
  - Applications in Astronomy
- **Present**

# Introduction

## Bayesian Statistics

Let  $\mathbf{x}$  be a latent field and  $\mathbf{y}$  an observable

- Prior density:  $p(\mathbf{y}|\mathbf{x})$
- Posterior density:  $p(\mathbf{x}|\mathbf{y})$
- Joint density:  $p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x}) \cdot p(\mathbf{y}|\mathbf{x}) = p(\mathbf{y}) \cdot p(\mathbf{x}|\mathbf{y})$
- Marginal posterior density:  $p(x_i|\mathbf{y}) = \int p(\mathbf{x}|\mathbf{y}) dx_{-i}$

# Introduction

## Latent Gaussian Models (LGMs)

- Bayesian Additive Models (BAMs)

- $y_i$  is assumed to belong to an exponential family with mean  $\mu_i$

- $y_i$  is linked to a structured additive predictor,  $\eta_i$ , via a link function  $g(\cdot)$ , such that  $g(\mu_i) = \eta_i$ , where

$$\eta_i = \mu + \sum_{j=1}^{n_f} f^{(j)}(\mathbf{u}_{ji}) + \sum_{k=1}^n \beta_k z_{ki} + \epsilon_i \quad (1)$$

- LGMs are a subset of BAMs, with a predictor as (1) and which assign a Gaussian prior to  $\mu$ ,  $\{f^{(j)}(\cdot)\}$ ,  $\{\beta_k\}$  and  $\{\epsilon_i\}$ .

- **Applications:** relaxation of regression models, dynamic models ( $u_t$ ), spatial models ( $u_s$ ), ...

# Introduction

## Latent Gaussian Models (LGMs)

- **Notation**

- $(\cdot|\cdot)$  - conditional density of its arguments
- $\mathbf{x}$  - all  $n$  Gaussian variables  $\{x_i\}$ ,  $\{f^{(i)}(\cdot)\}$ ,  $\{x_k\}$  and  $\{x_i\}$
- $(\mathbf{x}|\boldsymbol{\theta}_1)$  - is Gaussian with assumed zero mean, precision matrix  $\mathbf{Q}(\boldsymbol{\theta}_1)$ , and hyperparameters  $\boldsymbol{\theta}_1$
- $N(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$  - Gaussian density  $N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  at configuration  $\mathbf{x}$
- $(\mathbf{y}|\mathbf{x}, \boldsymbol{\theta}_2)$  - the distribution for the  $n_d$  observables  $\mathbf{y}$  (assumed conditionally independent given  $\mathbf{x}$  and  $\boldsymbol{\theta}_2$ )
- $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)^T$ , with  $\dim(\boldsymbol{\theta}) = m$



## Introduction

# Latent Gaussian Models (LGMs)

The posterior then reads:

- $(\mathbf{x}, \boldsymbol{\theta} | \mathbf{y}) \propto (\boldsymbol{\theta}) (\mathbf{x} | \boldsymbol{\theta}) \prod_i (y_i | x_i, \boldsymbol{\theta})$   
 $\propto (\boldsymbol{\theta}) |\mathbf{Q}(\boldsymbol{\theta})|^{1/2} \exp\left(-\frac{1}{2} \mathbf{x}^T \mathbf{Q}(\boldsymbol{\theta}) \mathbf{x} + \sum_i \log(y_i | x_i, \boldsymbol{\theta})\right)$

- **Properties** (satisfied by many LGMs but not all)

- Latent field  $\mathbf{x}$  admits conditional independence properties, making it a Gaussian Markov random field with a sparse precision matrix  $\mathbf{Q}(\boldsymbol{\theta})$
- The number of hyperparameters,  $m$ , is small ( $m \leq 6$ )

Both are usually required to produce fast inference

# Introduction

## Inference

- **Aim:** infer posterior marginals for  $(x_i|\mathbf{y})$ ,  $(\theta|\mathbf{y})$  and  $(\cdot_j|\mathbf{y})$
- **Possibilities:**
  - Markov Chain Monte Carlo
    - Poor performance when applied to LGMs
  - Deterministic approximations
    - Better computational cost

## A different approach

- The posterior marginals of interest can be written as

$$p(x_i | \mathbf{y}) = \int p(x_i | \boldsymbol{\theta}, \mathbf{y}) p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}$$

$$p(\boldsymbol{\theta}_{-j} | \mathbf{y}) = \int p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}_{-j}$$

- INLA (Rue et al. (2009)) uses this form to construct nested approximations

$$\tilde{p}(x_i | \mathbf{y}) = \int \tilde{p}(x_i | \boldsymbol{\theta}, \mathbf{y}) \tilde{p}(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}$$

$$\tilde{p}(\boldsymbol{\theta}_{-j} | \mathbf{y}) = \int \tilde{p}(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}_{-j}$$

where  $\tilde{p}(\cdot | \cdot)$  is an approximated density of its arguments and the integrations are performed numerically. The Laplace approximation of  $p(\boldsymbol{\theta} | \mathbf{y})$  is given by

$$\tilde{p}(\boldsymbol{\theta} | \mathbf{y}) \propto \frac{p(\mathbf{x}, \boldsymbol{\theta}, \mathbf{y})}{\tilde{p}_i(\mathbf{x} | \boldsymbol{\theta}, \mathbf{y})} \Big|_{\mathbf{x}=\mathbf{x}^*(\boldsymbol{\theta})} \quad (2)$$

1) Exploring  $\tilde{(\boldsymbol{\theta}|\mathbf{y})}$ 

a) locate mode of  $\tilde{(\boldsymbol{\theta}|\mathbf{y})}$ ,  $\boldsymbol{\theta}^*$ : using the difference between successive gradient vectors, approximate second derivatives of  $\log(\tilde{(\boldsymbol{\theta}|\mathbf{y})})$ ;

b) at  $\boldsymbol{\theta}^*$  compute the negative Hessian matrix  $\mathbf{H} > 0$  and let  $\boldsymbol{\Sigma} = \mathbf{H}^{-1}$ ; use standardized variables  $\mathbf{z}$  instead of  $\boldsymbol{\theta}$ , using the form

$$\boldsymbol{\theta}(\mathbf{z}) = \boldsymbol{\theta}^* + \mathbf{V} \boldsymbol{\Lambda}^{1/2} \mathbf{z}$$

c) explore  $\log(\tilde{(\boldsymbol{\theta}|\mathbf{y})})$ : start from the mode ( $\mathbf{z} = \mathbf{0}$ ); go in the positive direction of  $z_1$  with step  $\delta_z$ , while

$$\log(\tilde{(\boldsymbol{\theta}(\mathbf{0})|\mathbf{y}))} - \log(\tilde{(\boldsymbol{\theta}(\mathbf{z})|\mathbf{y}))} < \# \quad (3)$$

then switch direction; treat the remaining coordinates in the same way (fig. 1)

# INLA

## The Method

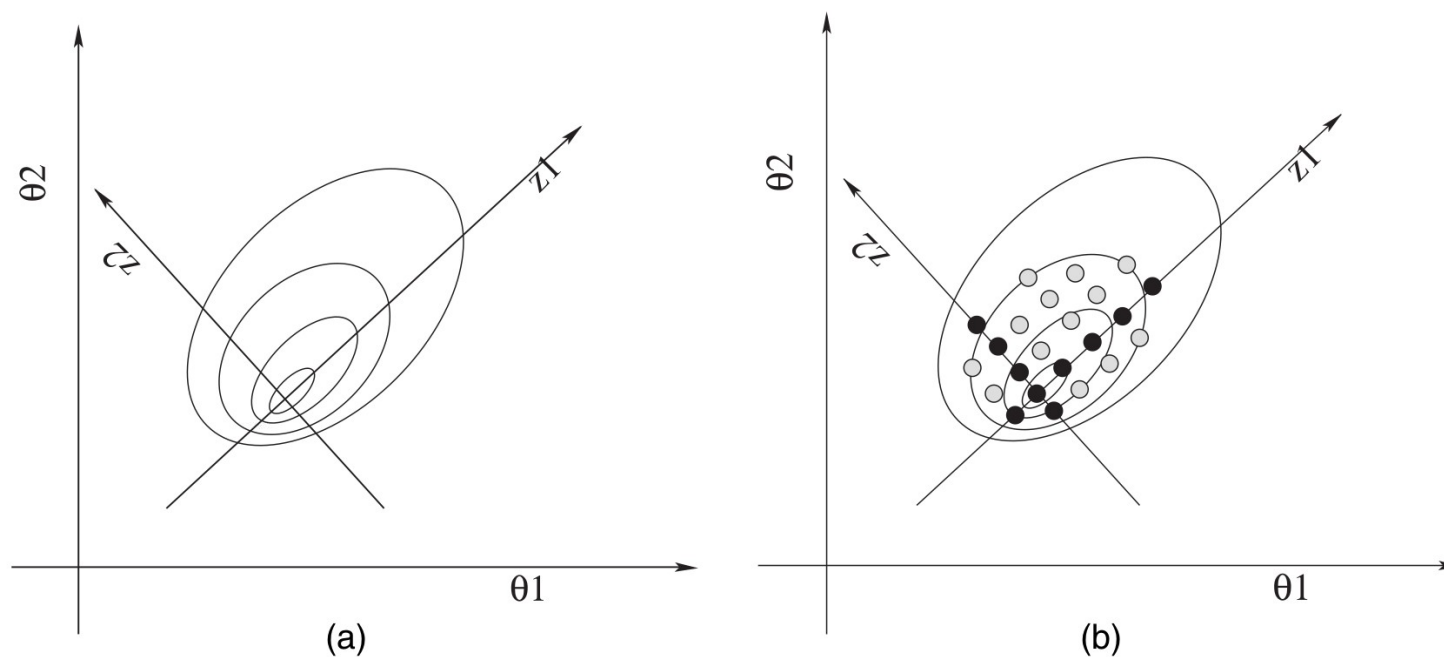


Fig. 1 - Illustration of the posterior marginal for  $\theta$ : in (a) the mode is located and the Hessian and co-ordinate system for  $\mathbf{z}$  are computed; in (b) each co-ordinate direction is explored ( $\bullet$ ) while  $\mathbf{z}$  prevails; new points ( $\circ$ ) are explored combining coordinates of ( $\bullet$ ) (source: Rue et al. (2009)).

d) use computed before to construct an interpolant to  $\log(\tilde{\boldsymbol{\theta}}|\mathbf{y})$  and compute marginals using numerical integration from this interpolant

## 2- Approximating $(x_i|\boldsymbol{\theta}, \mathbf{y})$

a) Approximate the modal configuration

$$\mathbf{x}_{-i}^*(x_i, \boldsymbol{\theta}) \approx \arg\max_{\mathbf{x}_{-i}} \tilde{\pi}_i(\mathbf{x}_{-i}|x_i) \quad (4)$$

b) Define a ROI around  $i$ ,  $\mathcal{S}_i(\boldsymbol{\theta})$ , for only those  $x_j$  ‘close’ to  $x_i$  should have an effect on its marginal;

c) Consider the Laplace approximation

$$\tilde{\pi}_i(x_i|\boldsymbol{\theta}, \mathbf{y}) \propto \frac{\tilde{\pi}(\mathbf{x}, \boldsymbol{\theta}, \mathbf{y})}{\tilde{\pi}_i(\mathbf{x}_{-i}|x_i, \boldsymbol{\theta}, \mathbf{y})} \Big|_{\mathbf{x}_{-i}=\mathbf{x}_{-i}^*(x_i, \boldsymbol{\theta})} \quad (5)$$

d) derive a simplified Laplace approximation  $\tilde{\gamma}_i(\mathbf{x}_i|\boldsymbol{\theta}, \mathbf{y})$  by doing a series expansion of  $\tilde{\gamma}_i(\mathbf{x}_i|\boldsymbol{\theta}, \mathbf{y})$  around  $\mathbf{x}_i = \boldsymbol{\mu}_i(\boldsymbol{\theta})$

e) expanding the log densities of both numerator and denominator in (5) around  $\mathbf{x}_i = \boldsymbol{\mu}_i(\boldsymbol{\theta})$ , we get

$$\log(\tilde{\gamma}_i(\mathbf{x}_i^s|\boldsymbol{\theta}, \mathbf{y})) = \text{const} - \frac{1}{2}(\mathbf{x}_i^s)^2 + \boldsymbol{\mu}_i^{(1)}(\boldsymbol{\theta}) \mathbf{x}_i^s + \frac{1}{6}(\mathbf{x}_i^s)^3 + \boldsymbol{\mu}_i^{(3)}(\boldsymbol{\theta}) + \dots \quad (6)$$

where

$$s_i^{(1)}(\boldsymbol{\theta}) = \frac{1}{2} \sum_{j \in \mathcal{N}_i} s_j^2(\boldsymbol{\theta}) \{1 - \rho_{ij}(\boldsymbol{\theta})\} d_j^{(3)}\{\mu_i(\boldsymbol{\theta}), \boldsymbol{\theta}\} s_j(\boldsymbol{\theta}) +_{ij}(\boldsymbol{\theta})$$

$$s_i^{(3)}(\boldsymbol{\theta}) = \sum_{j \in \mathcal{N}_i} d_j^{(3)}\{\mu_i(\boldsymbol{\theta}), \boldsymbol{\theta}\} \{s_j(\boldsymbol{\theta}) +_{ij}(\boldsymbol{\theta})\}^3$$

$$d_j^{(3)}(x_i, \boldsymbol{\theta}) = \frac{\partial^3}{\partial x_j^3} \log \{ p(y_j | x_j, \boldsymbol{\theta}) \} \Big|_{x_j = \mu_j(x_i | x_i)}$$

$$x_i^s = \frac{x_i - \mu_i(\boldsymbol{\theta})}{s_i(\boldsymbol{\theta})}$$



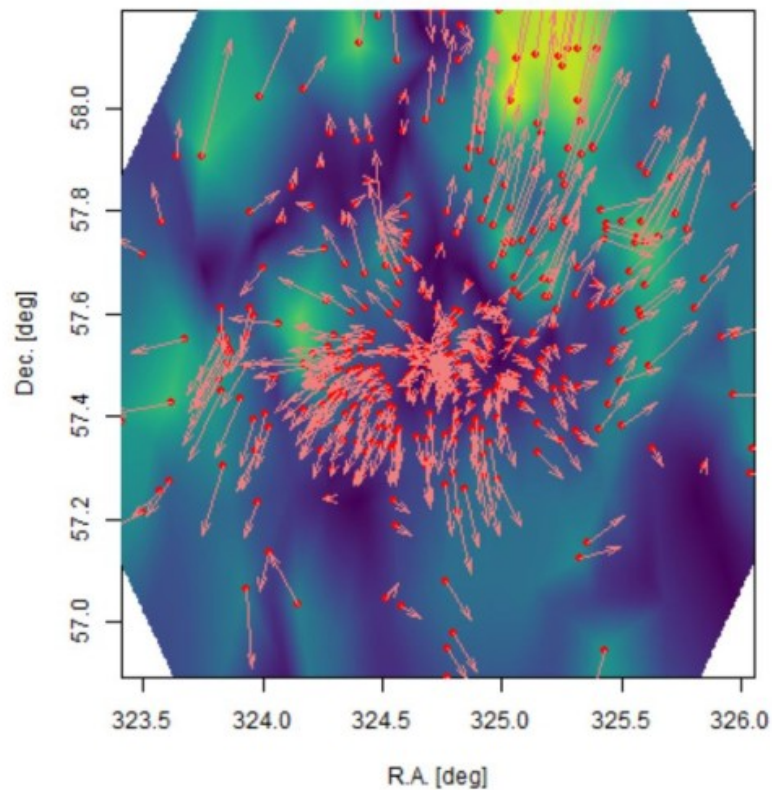
f) finally, fit a skew normal distribution of the form (7) to (6) so that the third derivative at the mode is  $\mu_i^{(3)}$ , the mean is  $\mu_i^{(1)}$  and the variance is 1.

$$\phi_{\mathcal{N}}(z) = \frac{2}{\omega} \phi\left(\frac{z - \xi}{\omega}\right) 1\left(+\frac{z - \xi}{\omega}\right) \quad (7)$$

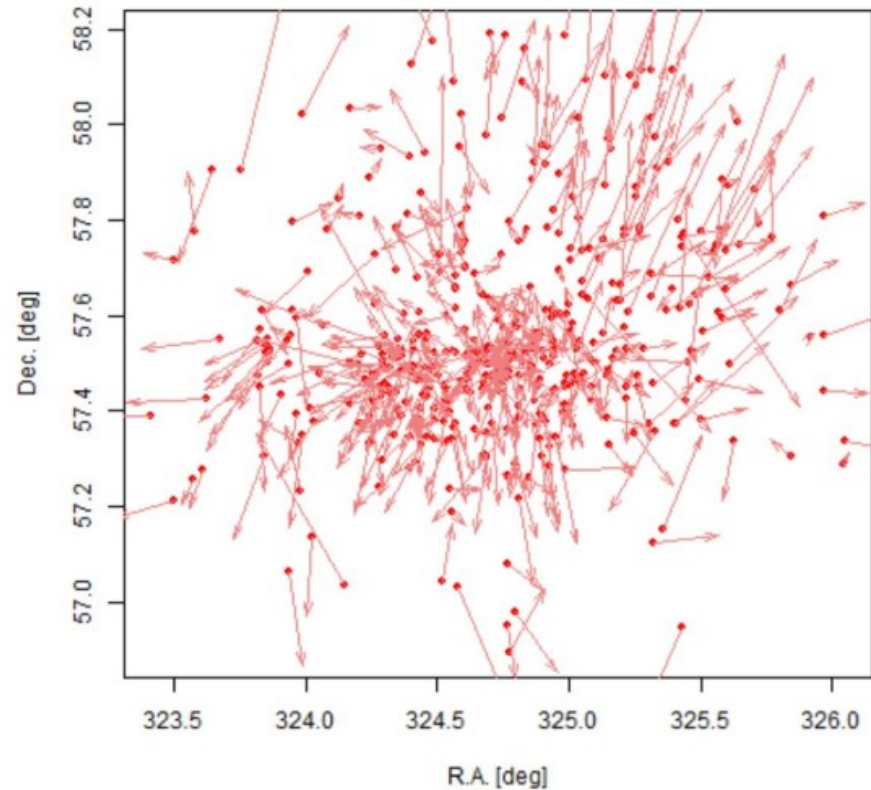
$\phi(\cdot)$  - density function  
 $1(\cdot)$  - distribution function  
 $a$  - skewness parameter  
 $\xi$  - location parameter  
 $\omega$  - scale parameter

# INLA

## Applications in Astronomy



IC 1396, inferred data

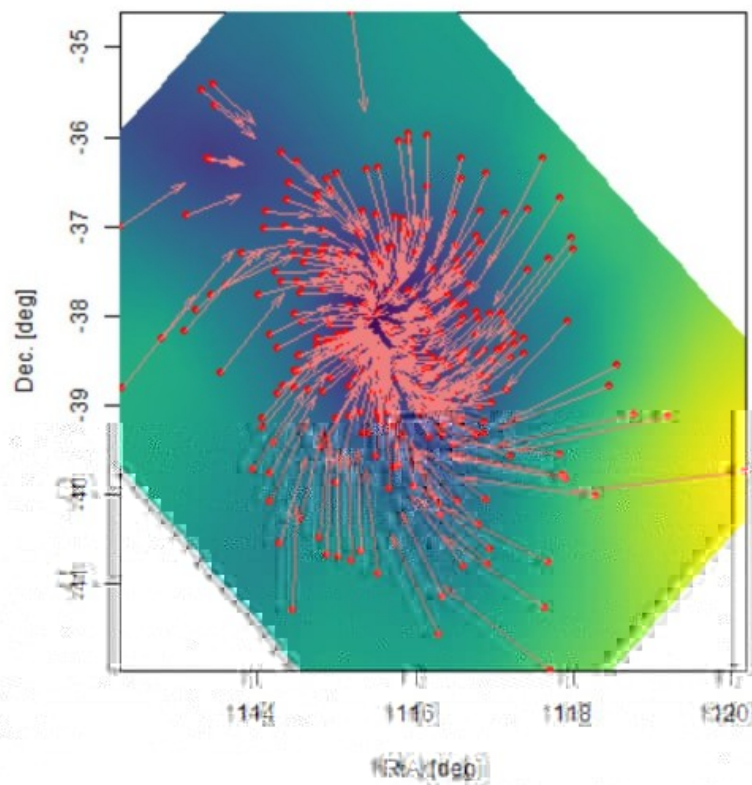


IC 1396, real data

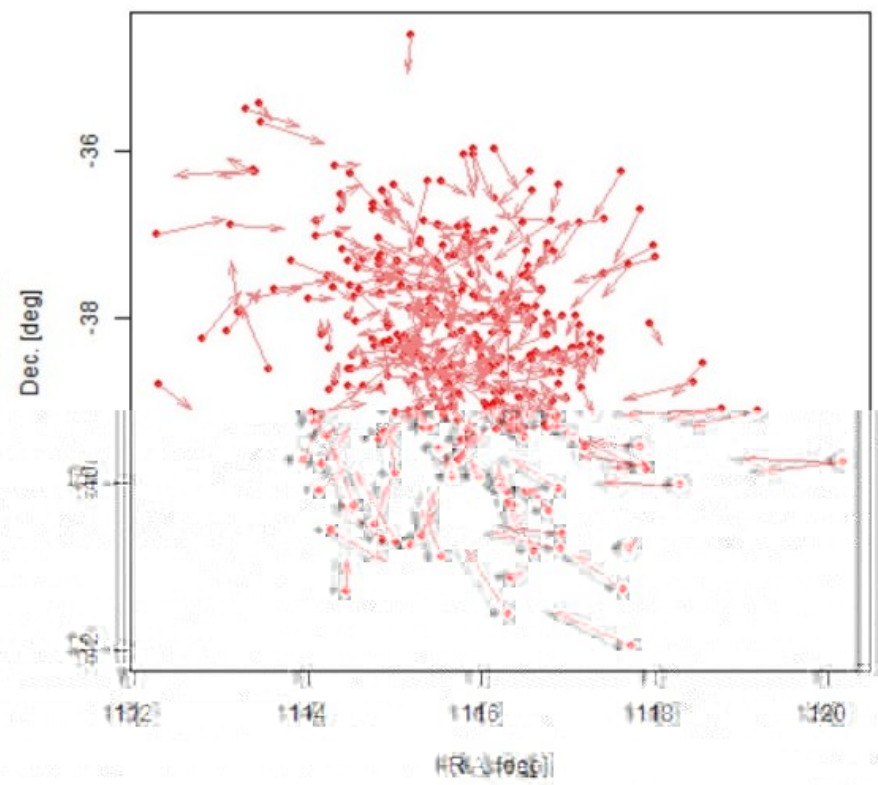
(source, Garcia et al. (2020))

# INLA

## Applications in Astronomy



NGC 2451A, inferred data

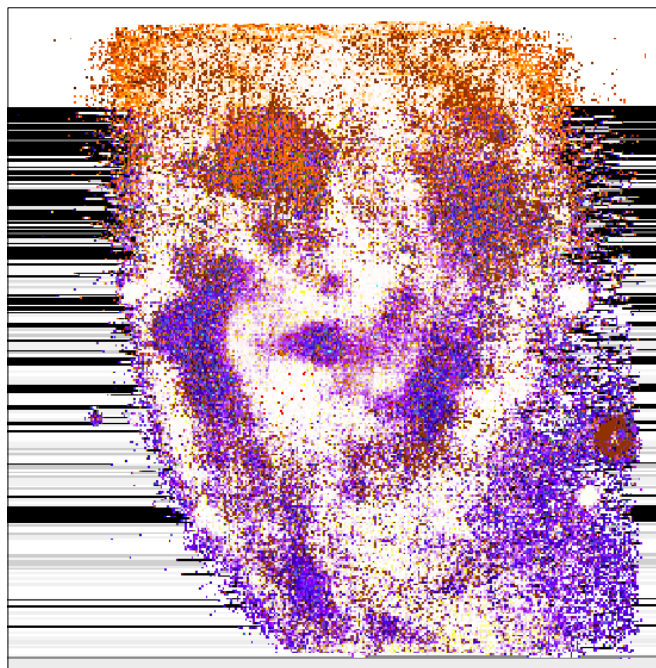


NGC 2451A, real data

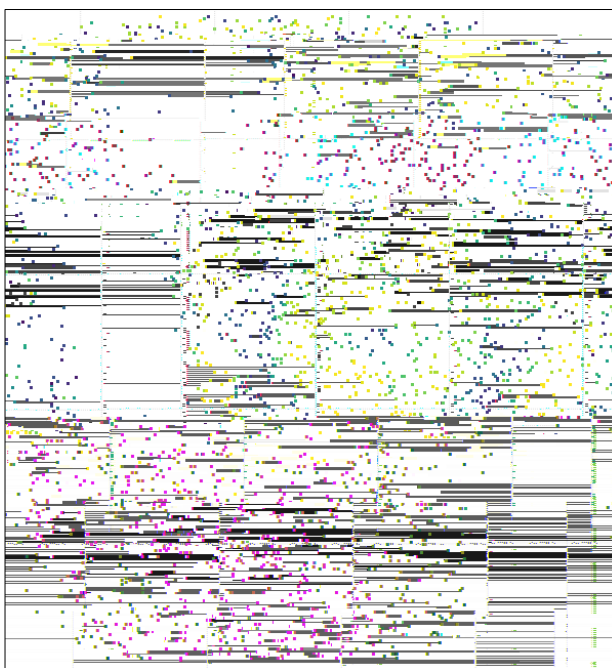
(source, Garcia et al. (2020))

# INLA

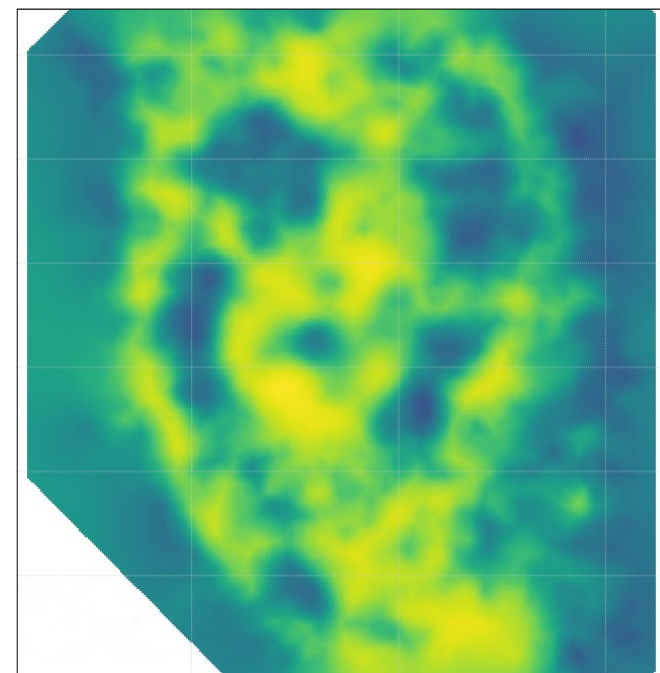
## Applications in Astronomy



ASASSN15db\_agel  
real data



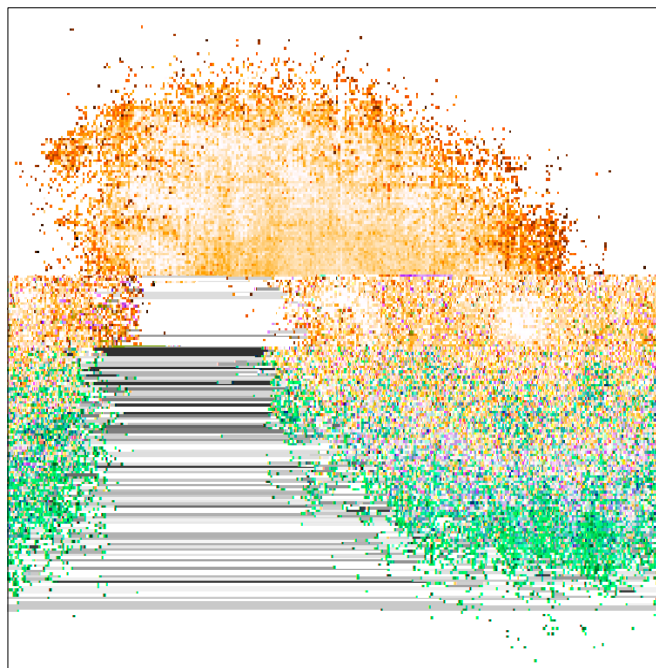
ASASSN15db\_agel  
5% sampling of real data



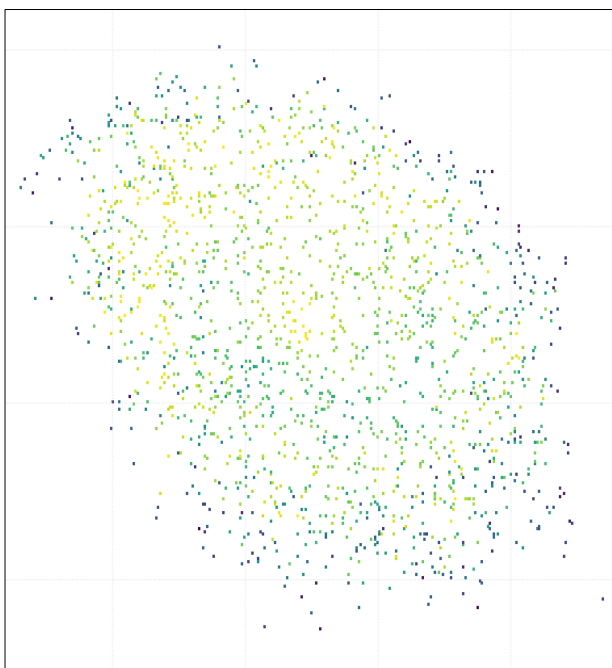
ASASSN15db\_agel  
INLA reconstruction

# INLA

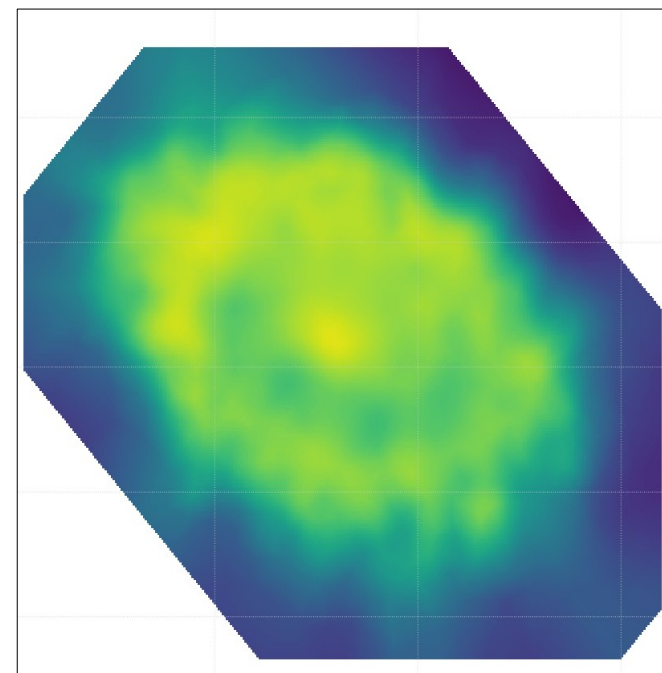
## Applications in Astronomy



PTF11qnr\_agel  
real data



PTF11qnr\_agel  
5% sampling of real data



PTF11qnr\_agel  
INLA reconstruction

- **INLA + Monte Carlo Radiative Transfer (MCRT)**

- 1) Generate low resolution simulations of radiative transfer using MC
- 2) Preprocess output files
- 3) Feed results as priors to INLA
- 4) Get high resolution posteriors in a fraction of the time

**Thank you**